

# Inequalities Between First and Second Order Moments for Continuous Probability Distribution



## S. R. Sharma, M. Gupta, Reetu Malhotra

Abstract: In this research article we obtained some inequalities between moments of 1st and 2nd order for a continuous distribution over the interval [x, y], when infimum and supremum of the continuous probability distribution is taken into consideration. These inequalities have shown improvement and are better than those exist in literature. Inequalities also obtained for continuous random variables which vary in [x, y] interval, such that the probability density function (pdf)  $\phi(t)$  become zero in  $[p, q] \subset [x, y]$ . The improvement in inequalities have been shown graphically. Here in this paper we deduced some existing inequalities by using the inequalities obtained in Theorem 2.1 and Theorem 2.2.

Key Words: Moments, Variance, Random Variable (R.V.), Infimum and Supremum

#### I. INTRODUCTION

If  $\mu_1^{\prime}$  be mean,  $\sigma^2$  be variance and  $\mu_2^{\prime}$  the 2nd order moment, the for R.V. in the interval [x, y], we have following:

$$\mu_1' = \int_x^y t\phi(t)dt \tag{1.1}$$

$$\sigma^2 = \int_{x}^{y} (t - \mu_1^{/})^2 \phi(t) dt$$
 and (1.2)

$$\mu_2' = \int_{x}^{y} t^2 \phi(t) dt \,. \tag{1.3}$$

Relationship between mean ( $\mu_1^{\prime}$ ), variance ( $\sigma^2$ ) and second order moment about origin ( $\mu_2^{\prime}$ ) is:

$$\mu_2' = \sigma^2 + \mu_1^{/2} \,. \tag{1.4}$$

Bounds of the variance  $(\sigma^2 = \mu_2' - \mu_1'^2)$  of a R.V., its extensions and applications have been found in literature.

#### Revised Manuscript Received on February 28, 2020.

\* Correspondence Author

S. R. Sharma, Chitkara University School of Engineering and Technology, Chitkara University, Himachal Pradesh (India) Email: sita.ram@chitkarauniversity.edu.in

M. Gupta, Chitkara University School of Engineering and Technology, Chitkara University, Himachal Pradesh (India) Email: madhu.gupta@chitkarauniversity.edu.in

**Reetu Malhotra\***, Chitkara University Institute of Engineering and Technology, Chitkara University, Punjab (India). Email: reetu.malhotra@chitkarauniversity.edu.in

© The Authors. Published by Blue Eyes Intelligence Engineering and Sciences Publication (BEIESP). This is an open access article under the CC-BY-NC-ND license <a href="http://creativecommons.org/licenses/by-nc-nd/4.0/">http://creativecommons.org/licenses/by-nc-nd/4.0/</a>

These inequalities have been discussed extensively over the period of time by B.C. Rennie, "On a class of Inequalities" (1963), D.S. Mitrinovic, "Analytic Inequalities" (1970), J.N. Kapur et al (1995), M.B. Benerjee et. al (1995), J.N. Kapur (1996), N.S. Barnett et al (1999), R. Bhatia and C. Davis in(2000), R Sharma et. al (2010). J.N. Kapur in (1996) obtained the following inequality when R.V. fall in the interval [x, y]:

$$0 \le \sigma^2 \le (y - \mu_1^{/})(\mu_1^{/} - x) \tag{1.5}$$

(1.5) can be written as

$$0 \le \sigma^2 + \left(\mu_1' - \frac{x+y}{2}\right)^2 \le \left(\frac{y-x}{2}\right)^2. \tag{1.6}$$

J.N. Kapur in (1995) has shown that the 2nd order moment  $\mu_2$  about origin lies between following maximum and minimum values:

$$\mu_1^{\prime 2} \le \mu_2^{\prime} \le (x+y)\mu_1^{\prime} - xy$$
 (1.7)

Inequality (1.5) and (1.7) are equivalent.

The plot of inequality (1.7) shows that point ( $\mu_1', \mu_2'$ ) lies in  $\mu_1', \mu_2'$  - plane region between quadratic and linear curve of moments as shown in figure (1.1).

Such type of basic inequalities their extensions, refinements was studied by different mathematicians. Alternative proofs for these inequalities also given by several authors. In particular Sharma et. al (2012) had given proofs for the inequalities involving moments of discrete uniform distributions.

$$\mu_2' \ge \frac{x^2 + y^2}{N} + \frac{N}{N - 2} \left[ \mu_1' - \frac{x + y}{N} \right]^2$$

This inequality shows refinement in inequality (1.7) given by J N Kapur in (1996). Here in this paper first two moments  $\mu_1$  and  $\mu_2$  of a R.V. which is continuous and which lie in interval [x, y] are considered. We derive some inequalities which involve 1st and 2nd order moments about origin for a continuous probability distribution in the presence of minimum and maximum values of the pdf when R.V. lies in the interval [x, y]. A bound on variance of a R.V. which lies in the interval [x, y] was discussed by Kapur et. al, (1996) and is given by inequalities (1.5),(1.6) and (1.8).



# Inequalities Between First and Second Order Moments for Continuous Probability Distribution

Here we have obtained some refinements in (1.7) subject to the condition that pdf is taken into consideration (both minimum and maximum values) and also discussed related refinements and extensions. The main results (Theorem 2.1 and Theorem 2.2) gives bound for 2nd order moment and 1st order moment subject to above condition that also shows refinement in (1.7).

Sharma et. al. (2015) proved following inequalities (1.8) and (1.9) below for the case when, mean is  $\mu_1^{\prime}$  and standard deviation is  $\sigma$  for R.V. "t"  $\phi(t)$  is pdf defined in [x, y], m be the infimum and M is supremum of the function  $\phi(t)$  in the interval [x, y], then

$$\frac{m(y-x)^3}{12} \le \sigma^2 + \left(\mu_1' - \frac{x+y}{2}\right)^2 \le \left(\frac{y-x}{2}\right)^2 - \frac{m(y-x)^3}{6}$$
 and

$$\left(\frac{y-x}{2}\right)^{2} - \frac{M(y-x)^{3}}{6} \le \sigma^{2} + \left(\mu_{1}' - \frac{x+y}{2}\right)^{2} \le M \frac{(y-x)^{3}}{12}$$
(1.9)

In Theorem 2.3 it has been shown that above bounds can be deduced by using Theorem 2.1 and Theorem 2.2 below.

We also discuss the results i.e. in theorem 2.1 and 2.2 geometrically. Theorem 2.4 shows how the region in figure 1.1 for inequality (1.7) reduces to new region for new inequality and shown by figure 1.2.

### II. MAIN RESULTS

# **THEOREM 2.1**

Let  $\mu_1^{\prime}$  and  $\mu_2^{\prime}$  be the moments,  $\phi(t)$  be pdf and "t" a continuous R.V. whose pdf. is defined in the interval [x, y]. If m be the greatest lower bound for pdf  $\phi(t)$  in [x, y] then we must have.

$$(x+y)\mu_1' - \frac{(x+y)^2}{4} + \frac{m(y-x)^3}{12} \le \mu_2' \le (x+y)\mu_1' - xy - \frac{m(y-x)^3}{6}$$
(1.10)

# **PROOF**

Consider

$$\int_{x}^{y} [t^{2} - (x+y)t + xy]\phi(t)dt = \int_{x}^{y} (t-y)(t-x)\phi(t)dt$$

Since t lies in the interval [x, y] therefore,

$$(t-x)(t-y) \le 0$$

$$\Rightarrow \int_{x}^{y} [t^{2} - (x+y)t + xy]\phi(t)dt \le m \int_{x}^{y} (t-y)(t-x)dt$$

$$\Rightarrow \mu_2' \le (t+y)\mu_1' - ty - \frac{m(y-x)^3}{6}. \tag{1.11}$$

Now 
$$t^2 - (t + y)t + ty = \left(t - \frac{x + y}{2}\right)^2 - \left(\frac{y - x}{2}\right)^2$$

$$\Rightarrow \int_{x}^{y} \left( t^{2} - (x+y)t + \frac{(y+x)^{2}}{4} \right) \phi(t) dt =$$

$$\int_{x}^{y} \left( t - \frac{x+y}{2} \right)^{2} \phi(t) dt$$
(1.12)

Using properties of pdf  $\phi(t)$  over the limits x to y in (1.12) we have

$$\Rightarrow \mu_2^{/} \ge (t+y)\mu_1^{/} - \left(\frac{x+y}{2}\right)^2 + m\frac{(y-x)^3}{12} . (1.13)$$

Inequality (1.10) now follows from inequality (1.11) and (1.13).

By using algebraic operations and techniques of derivatives for maxima and minima we can easily show that

$$\mu_1^{/2} \le (x+y)\mu_1^{/} - \frac{(x+y)^2}{4} + \frac{m(y-x)^3}{12} \le \mu_2^{/} \le$$

$$(x+y)\mu_1' - xy - \frac{m(y-x)^3}{6} \le (x+y)\mu_1' - xy$$

This shows refinement in (1.7).

#### **THEOREM 2.2**

Let  $\mu_1^{\prime}$  and  $\mu_2^{\prime}$  be moments of a continuous R.V. "t" whose pdf  $\phi(t)$  is defined in [x, y]. If M be the least upper bound of the function  $\phi(t)$  in [x, y] then we must have,

$$(x+y)\mu_{1}' - xy - \frac{M(y-x)^{3}}{6} \le \mu_{2}' \le$$

$$(x+y)\mu_{1}' - \left(\frac{x+y}{2}\right)^{2} + \frac{M(y-x)^{3}}{12}$$

$$(1.14)$$

# ROOF

Consider

$$\int_{x}^{y} [t^{2} - (x+y)t + xy]\phi(t)dt = \int_{x}^{y} (t-y)(t-x)\phi(t)dt$$

Since t  $\mathcal{E}[x, y]$  hence we have:

$$(t - x) (t - y) < 0$$

$$\Rightarrow \int_{x}^{y} [t^{2} - (x+y)t + xy]\phi(t)dt \ge M \int_{x}^{y} (t-y)(t-x)dt$$

$$\Rightarrow \mu_2' \ge (t+y)\mu_1' - xy - \frac{M(y-x)^3}{6}. \tag{1.15}$$

Now

$$t^{2} - (x + y) t + xy = \left(t - \frac{x + y}{2}\right)^{2} - \left(\frac{y - x}{2}\right)^{2}$$

$$\Rightarrow \int_{x}^{y} \left(t^{2} - (x + y)t + \left(\frac{y + x}{2}\right)^{2}\right) \phi(t)dt =$$

$$\int_{x}^{y} \left(t - \frac{x + y}{2}\right)^{2} \phi(t)dt$$
(1.16)

On simplifying (1.16) we have





$$\mu_2' \le (x+y)\mu_1' - \left(\frac{y+x}{2}\right)^2 + \frac{M(y-x)^3}{12}$$
 (1.17)

Inequality (1.14) now follows from inequality (1.15) and (1.17).

Similarly (1.14) is refinement in (1.7) and is given as

$$\mu_1^{/2} \le (x+y)\mu_1^{/} - xy - \frac{M(y-x)^3}{6} \le \mu_2^{/} \le$$

$$(x+y)\mu_1' - \left(\frac{x+y}{2}\right)^2 + \frac{M(y-x)^3}{12} \le (x+y)\mu_1' - xy$$

If t be a R.V. in [x, y] and  $\phi(t)$  be pdf then inequalities (1.8) and (1.9) can be deduced from (1.10) and (1.14) respectively.

We know that  $\mu_1^{\prime}$ ,  $\mu_2^{\prime}$  and  $\sigma$  are related by the following equation:

$$\mu_2^{\prime} = \sigma^2 + \mu_1^{\prime} 2 \quad . \tag{1.18}$$

Substituting value of  $\mu_2^{\prime}$  from (1.18) in inequality (1.10), we get inequality (1.8). Similarly substituting the value of  $\mu_2$ from (1.18) in inequality (1.14), we get inequality (1.9).

#### **THEOREM 2.4**

If t be a R.V. which lies in [x,y], with pdf  $\phi(t)$ , such that  $\phi(t)$ vanishes in  $[p, q] \subset [x, y]$ , we must have,

$$(p+q) \ \mu_1' - pq \le \ \mu_2' \le (x+y) \ \mu_1' - xy. \tag{1.19}$$

From inequality (1.19) it is obvious that point  $(\mu_1^{\prime}, \mu_2^{\prime})$  in

 $\mu_1'$   $\mu_2'$  - plane lies on or below the curve

 $\mu_2' = (x + y) \ \mu_1' - xy$  and on or above the curve

 $\mu_2^{\prime} = (p + q) \mu_1^{\prime} - pq.$ 

This is shown in figure (1.2).

#### **PROOF**

Since  $(t-x)(t-y) \le 0$  for t in [x, y] then,  $t^2 \le (x + y) t - xy$ 

(1.20)If 't' lies outside i.e.  $t \notin [p, q]$  then,

$$(t-p) (t-q) \ge 0$$
  
 $t^2 \ge (p+q) t - pq$  (1.21)

Multiply (1.20) with  $\phi(t)$  the pdf and integrating with R.V. t between limits t = x and t = y we get maxima for inequality (1.19). Similarly, on multiply (1.21) by pdf  $\phi(t)$  and on integrating minima for inequality (1.19).

When  $\phi(t)$  vanishes in [p, q]  $\subset$  [x, y] the point  $(\mu_1, \mu_2)$ lies on or below the line joining points E and F. Further the point  $(\mu'_1, \mu'_2)$  lies on or above the line joining points C and D. Figure (1.2) shows that inequality (1.19) is refinement of inequality (1.7) shown by Figure (1.1).

### III. CONCLUSIONS

Theorem 2.1 and Theorem 2.2 shows that the results obtained are better than inequality (1.7). Theorem 2.3 proved that inequality (1.8) and (1.9) obtained by Sharma et..al. in (2015)

can be deduced from inequalities obtained in theorem2.1 and 2.2 by using the relation  $\mu_2^{\prime} = \sigma^2 + \mu_1^{\prime}^2$ . Theorem 2.4 proved that inequality (1.19) is better than (1.7) when pdf  $\phi$  (t) assume zero value in the interval [p,q] contained in [x, y]. Betterment in inequality (1.19) explained geometrically and it shows that shaded region given by inequality (1.19) in figure 1.2 is more compact than in figure 1.1 for inequality (1.7).

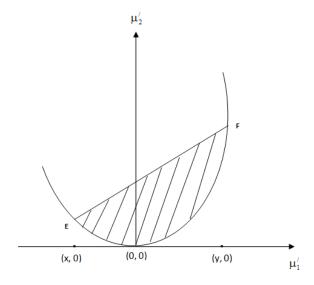


Figure: 1.1  $E \leftrightarrow (x, x^2)$  and  $F \leftrightarrow (y, y^2)$ .

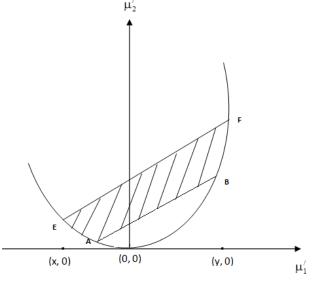


Figure: 1.2  $A \leftrightarrow (p, p^2)$   $B \leftrightarrow (q, q^2)$ 

$$E \leftrightarrow (x, x^2)$$
 and  $F \leftrightarrow (y, y^2)$ 

# REFERENCES

- 1. J.N. Kapur and Anju Rani, Testing the consistency of given values of a set of moments of a probability distribution, J. Bihar Math. Soc 16 (1995), 51-63.
- M.B. Benerjee and R.G.Shandil, A Theorem on Mean and Standard Deviation of statistical Variate, Ganita, 47, 21-23, (1995).



# Inequalities Between First and Second Order Moments for Continuous Probability Distribution

- N.S. Barnett and S. S. Drogomir, Some elementary Inequalities for the Expectation and Variance of a Random Variable whose pdf is defined on a finite interval, RGMIA Res. Rep. Coll, 2(7) (1999).
- S.R. Sharma et al, Some Elementary Inequalities Between Mean and Standard Deviation, Mathematical Journal of Interdisciplinary Sciences 4(1), (2015) 23–28.
- D.S..Mitrinovic, Analytic Inequalities, Springer Verlag-New York, (1970).
- B.C.Rennie, On a class of Inequalities, J. Austral. Math. Soc. 3, 442-448, (1963).
- S.R. Sharma and R. Sharma, On Inequalities Involving Moments of Discrete Uniform Distributions. Mathematical Journal of Interdisciplinary Sciences (MJIS),Vol.(1),45-56(2012).http://dx.doi.org/10.15415/mjis.2012.110 03
- 8. J.N.Kapur, On Inequalities Between Moments of Probability Distribution, Ganita, Vol. 47, No. 2 pp 37-41, (1996).
- 9. N.S. Barnett and S.S. Drogomir. Some Elementary Inequalities for the Expectation and Variance of a Random Variable whose pdf is defined on a finite interval. RGMIARes. Rep. Coll, 2(7), Art(1), (1999).
- N.S.Barnett, P.Cerone, S.S. Drogomir and J. Roumeliotis. Some Inequalities for the Expectation and Variance of a Random Variate whose pdf is n-times differentiable. JIPAM, Vol. 1(2), Art (21) (2000).
- R Bhatia, C. Davis, A Better Bound on the Variance. Amer. Math. Monthly, 107, 353-357(2000). http://dx.doi.org/10.2307/2589180.
- Edward Neuman, On Some Means Derived From The Schwab–Borchardt Mean II Journal of Mathematical Inequalities, Volume 8, and Number 2 359–368, (2014).

#### **AUTHORS PROFILE**



**Dr. S R Sharma**, is M. Sc., M. Phil, Ph. D, and is Associate Professor in Department of Applied Sciences, Chitkara University, and Himachal Pradesh, India since August, 2008. He completed his B. Sc. M.Sc. M. Phil and Ph. D from Himachal Pradesh University, Shimla. He has over 15 years of experience in Academic and Research. He

has published his work in various National and International Journals. His areas of interests include Inequalities, Applied Mathematics, Fractal Geometry, Elasticity and Engineering Education.



**Dr. M. Gupta**, (M. Phil, Ph. D), is Associate Professor in Department of Applied Sciences, Chitkara University, Himachal Pradesh, India since January 2010. She holds her M. Phil and Ph. D from Himachal Pradesh University, Shimla. She has over 10 years of experience in Academic

and Research. She has published her work in various National and International Journals. Her areas of interests include Inequalities, Applied Mathematics, Engineering Education and Statistics. She is a life member of ISTE and HSCA.



**Dr. Reetu Malhotra**, (M. Phil, Ph. D), is Associate Professor in Department of Applied Sciences, Chitkara University, Punjab, India since July, 2008. She has 15 years of experience in Academics and Research. Her areas of interests include Reliability Theory, Probability,

Statistics and Operations Research



Journal Website: www.ijitee.org